

Large Scale Longitudinal Experiments: Estimation and Inference

October 30, 2024

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Apoorva Lal

Paper (joint with Alex Fischer, Matthew Wardrop): [arXiv:2410.09952](https://arxiv.org/abs/2410.09952)

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Why study panel data in recommender systems?

- We often have longitudinal data for users/devices/... : $\{\mathbf{Z}_{i,t}\}_{i=1,t=1}^{N,T}$
 - In typical experiments, we *flatten the time dimension* and compute a difference in means in the post-treatment window (optionally adjusting for average outcome in pre-treatment window)
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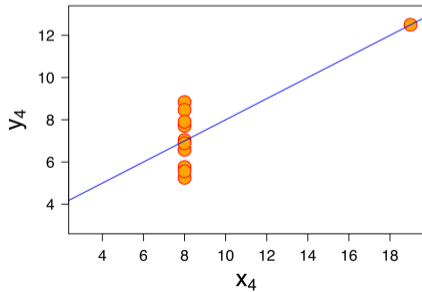
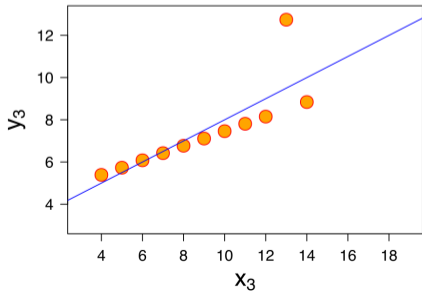
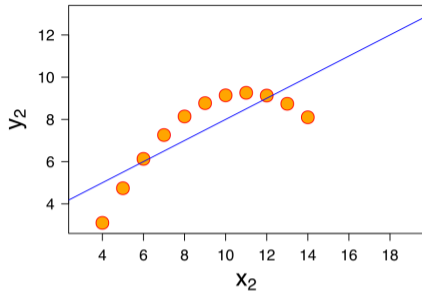
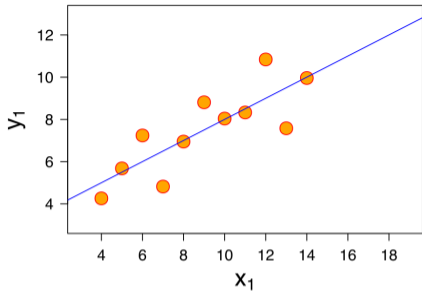
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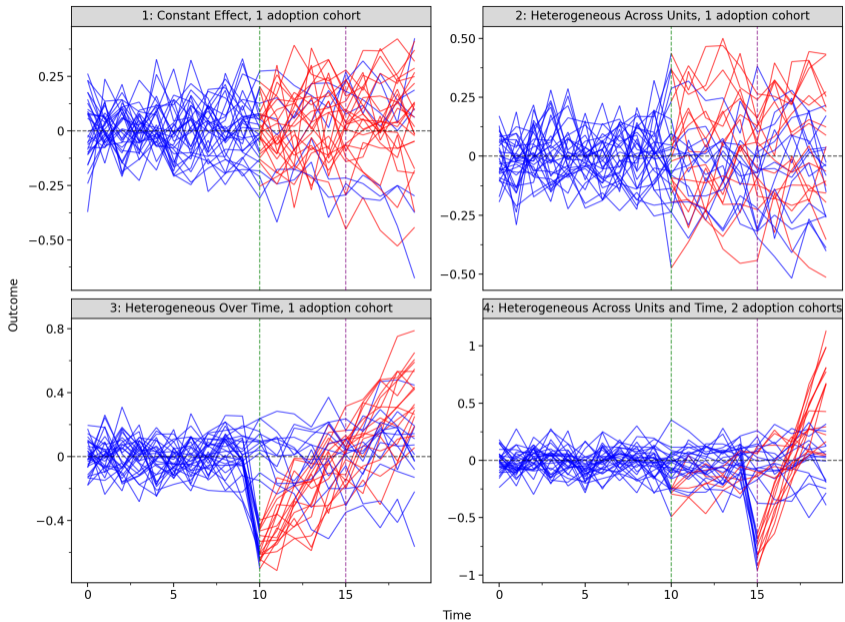
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 - **this project:** propose scalable panel data estimators that help identify these
 - temporal and cohort-level granularity is informative and important - don't flatten them with a T-test
- 



Raw outcomes for 30 units from four DGPs
Cross-sectional ATE=0 for all four



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Compressing Least Squares

$$\hat{\beta} = \left(\underbrace{\mathbf{X}^\top}_{P \times N} \underbrace{\mathbf{X}}_{N \times P} \right)^{-1} \underbrace{\mathbf{X}^\top}_{P \times N} \underbrace{\mathbf{y}}_{N \times 1}$$

$$\mathbb{V}[\hat{\beta}] = (\mathbf{X}^\top \mathbf{X})^{-1} \mathbf{X}^\top \Omega \mathbf{X} (\mathbf{X}^\top \mathbf{X})^{-1}$$



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- Suppose \mathbf{X} takes on values in \mathcal{X} with finite cardinality C
- Then, we compute sufficient statistics by strata, run $\tilde{\mathbf{y}}/\tilde{\mathbf{n}} \sim \tilde{\mathbf{X}}\beta + \varepsilon$: C observations



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Table 1: Example dataset and its compressed versions.

(a)	(b)	(c)	(d)
\mathbf{M} \mathbf{y}	$\tilde{\mathbf{M}}$ $\tilde{\mathbf{y}}$ $\tilde{\mathbf{n}}$	$\bar{\mathbf{M}}$ $\bar{\mathbf{y}}$ $\bar{\mathbf{n}}$	$\tilde{\mathbf{M}}$ $\tilde{\mathbf{y}}'$ $\tilde{\mathbf{y}}''$ $\tilde{\mathbf{n}}$
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A 1	A 2 1	B 3.5 2	B 7 25 2
A 2	B 3 1	C 5 1	C 5 25 1
B 3	B 4 1		
B 4	C 5 1		
C 5			

(a) Uncompressed data. (b) f-weights: (\mathbf{y}, \mathbf{M}) -compressed records.

(c) Groups: (\mathbf{M}) -compressed records. (d) Sufficient Statistics: (\mathbf{M}) -compressed records.

ABLaze is built around this: Wong et al(2021)

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$$\hat{\mathbb{V}}(\beta) = \overbrace{\left(\tilde{\mathbf{X}}^\top \text{diag}(\tilde{\mathbf{n}}) \tilde{\mathbf{X}} \right)^{-1}}^{\text{Bread}} \overbrace{\tilde{\mathbf{X}}^\top \text{diag} \left(\sum_j \underbrace{(\tilde{\mathbf{y}}_j^2 \tilde{\mathbf{n}}_j - 2\tilde{\mathbf{y}}_j \tilde{\mathbf{y}}_j' + \tilde{\mathbf{y}}_j'')}_{\text{RSS in } j} \right)}^{\text{Meat}} \overbrace{\left(\tilde{\mathbf{X}}^\top \text{diag}(\tilde{\mathbf{n}}) \tilde{\mathbf{X}} \right)^{-1}}^{\text{Bread}}$$

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Building Blocks

- Base untreated potential outcome : $Y_{it}^0 = \alpha_i + \gamma_t + \epsilon_{it}$
- Treated potential outcome under static, constant effects: $Y_{it}^1 = Y_{it}^0 + \tau$



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- Treated potential outcome under unit heterogeneity:
 $Y_{it}^1 = Y_{it}^0 + \tau_i$
- Treated potential outcome under time heterogeneity:
 $Y_{it}^1 = Y_{it}^0 + \sum_{k \geq 0} \tau_k 1\{t - T_i = k\}$



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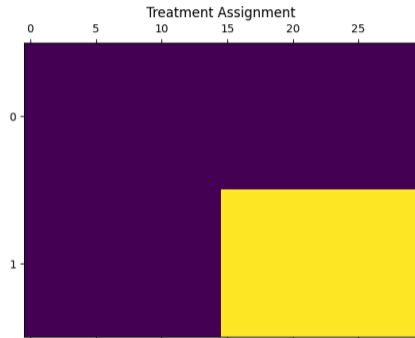
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- Observed outcome:

$$Y_{it} = Y_{it}^0(1 - W_{it}) + Y_{it}^1 W_{it}$$

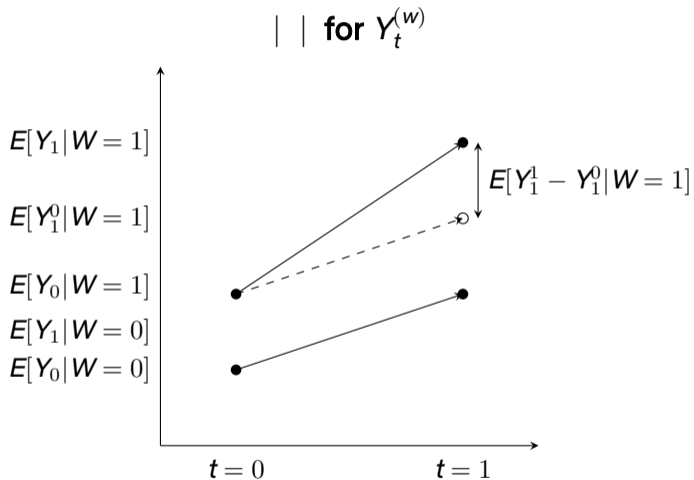
- Under random assignment or parallel trends, we can form estimates of Y_{it}^0 and construct estimates of (reasonable averages of) τ_{it}



When assignment time is endogenous, need generalized propensity score Arkhangelsky et al (2024)

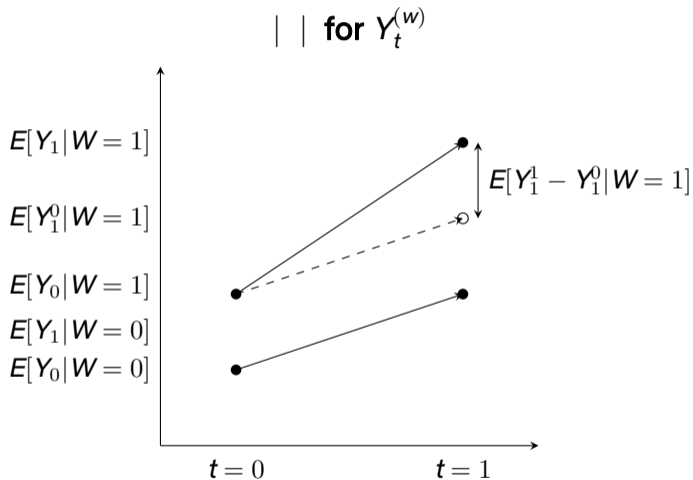
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Parallel Trends



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Parallel Trends



Easier to satisfy in experiments since $\mathbb{E}[Y_0 | W = 1] = \mathbb{E}[Y_0 | W = 0]$ under random assignment

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Regressions we'd like to run

$$\bar{Y}_{i,t>T_0} = \alpha + \tau W_i + \varepsilon_i$$

$$\bar{Y}_{i,t>T_0} = \alpha + \tau W_i + \beta \bar{Y}_{i,t<T_0} + \varepsilon_i$$

Diff in Means

CUPED

- $Z_{it} := \mathbb{1}\{\text{argmin}\{t : W_{it} = 1\} - t = k\}$ (treatment indicator \times event-time)



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Static TWFE

$$Y_{it} = \alpha_j + \gamma_t + \sum_{k=0, k \neq -1}^T \tau_k Z_{it}^k + \varepsilon_{it}$$

Dynamic TWFE (Event Study)

$$Y_{it} = \alpha + \gamma_t + \sum_{k \geq 0}^T \tau_k Z_{it}^k + \varepsilon_{it}$$

Dynamic DiM

$$Y_{it} = \alpha_j + \gamma_t + \sum_{c=2}^C \sum_{k=0, k \neq -1}^T \mathbb{1}\{G_i = c\} \tau_{kc} Z_{it}^k + \varepsilon_{it}$$

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- W_i, α_j, γ_t jointly identify a single observation - cannot compress
- Unit intercepts α_j : millions of distinct values

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Trick 1: Partialling out - the within estimator

- We don't inherently care about α_i, γ_t ; they are nuisance parameters
- Partial them out (i.e. kick them out of $\mathbf{X}'\mathbf{X}$)
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 - specialized methods for high dimensional categorical covariates: Method of alternating projections / Kaczmarz method (implemented in `areg/reghdfe/fixest/felm/pyfixest/...`)
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- Residualise RHS $W_{it} - \bar{W}_{i,\cdot} - \bar{W}_{\cdot,t} =: \ddot{W}_{it}$
- Obviates the need to adjust for time-invariant member characteristics (all absorbed in FEs)

$$Y_{it} = \tau \ddot{W}_{it} + \varepsilon_{it}$$

This regression *can* be compressed. However, we don't have an equivalent representation for dynamic regressions (event studies, disaggregated event studies), etc.



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Trick 2: Mundlak (1978) + Wooldridge (2021) Trick

$$Y_{it} = \alpha_j + \gamma_t + \tau W_{it} + \varepsilon_{it}$$

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Mundlak-ed Static TWFE



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Disagg Dynamic TWFE

$$Y_{it} = \underbrace{\delta + \sum_{c=1}^C \psi_c \mathbb{1}\{D_i = k\}}_{\text{Cohort Dummies}} + \underbrace{\sum_{k=1}^T \phi_t \mathbb{1}\{t = k\}}_{\text{Time Dummies}} + \underbrace{\sum_{c=1}^C \sum_{k=1}^T \tau_{kc} \mathbb{1}\{D_i = c\} \mathbb{1}\{t = k\}}_{\text{Cohort} \times \text{Time interactions}} + \varepsilon_{it}$$

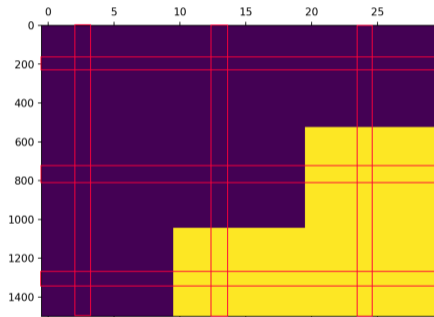


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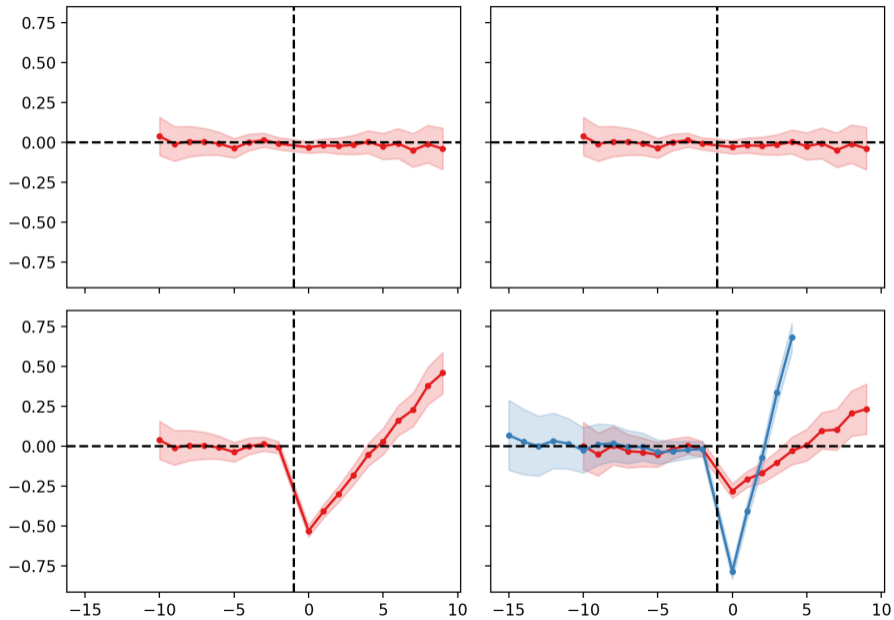
Design Matrix Dimensions

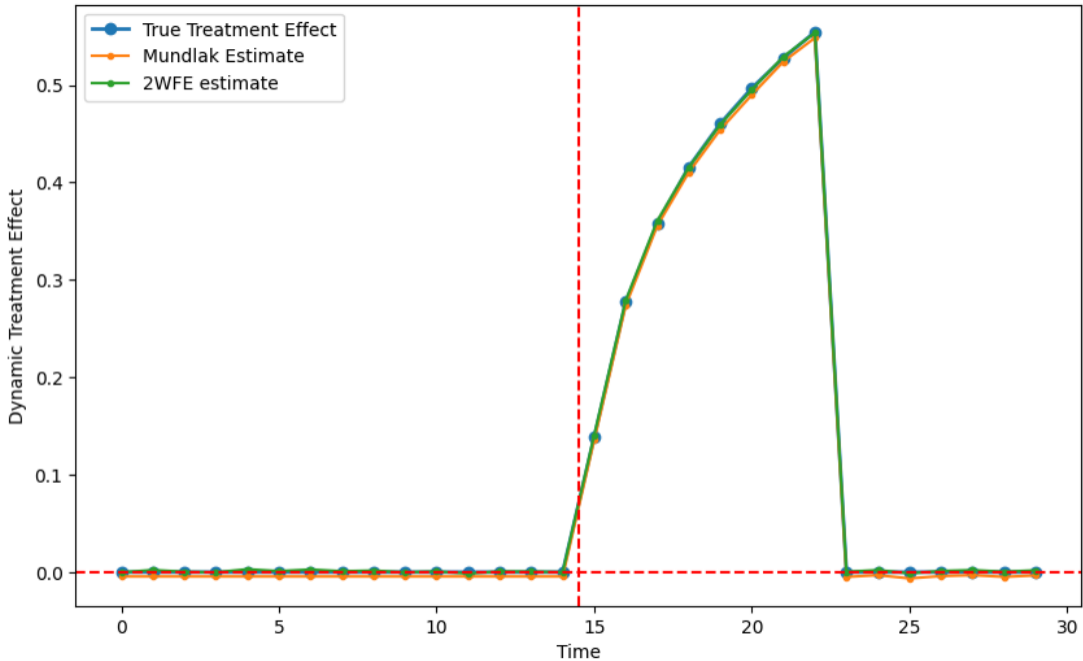
	(1) Standard	M	(2) Mundlak	\tilde{M}
Static	$Y_{it} = \alpha_i + \gamma_t + \tau W_{it} + \varepsilon_{it}$	NT	$Y_{it} = \alpha + \tau W_{it} + \psi \bar{W}_{i.} + \phi \bar{W}_{.t} + \varepsilon_{it}$	$2+(C+1)$
Dyn	$Y_{it} = \alpha_i + \gamma_t + \sum_{k \neq -1} \tau_k Z_{it}^k + \varepsilon_{it}$	NT	$Y_{it} = \alpha + \psi D_i + \sum_{k=1}^T \phi_t 1_{t=k} + \sum_{k=1}^T \tau_k D_i 1_{t=k} + \varepsilon_{it}$	2T
Dyn+Stagg	$Y_{it} = \alpha_i + \gamma_t + \sum_{c=1}^C \sum_{k \neq -1} \tau_{kc} 1_{G_i=c} Z_{it}^k + \varepsilon_{it}$	NT	$Y_{it} = \alpha + \sum_{c=1}^C \psi_c 1_{D_i=c} + \sum_{k=1}^T \phi_t 1_{t=k} + \sum_{c=1}^C \sum_{k=1}^T \tau_{kc} 1_{D_i=c} 1_{t=k} + \varepsilon_{it}$	CT

- N units, T time periods
- M, \tilde{M} : number of required rows in design matrix
- C unique adoption cohorts (including control)
- $2 + (C + 1) = 4$ obs in standard A/B test
- Cluster-robust inference with cluster-bootstrap

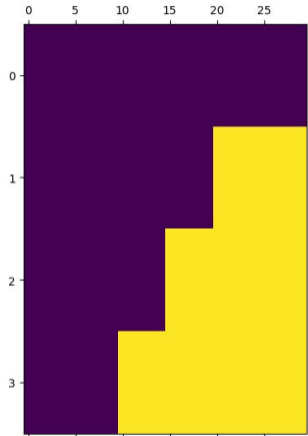
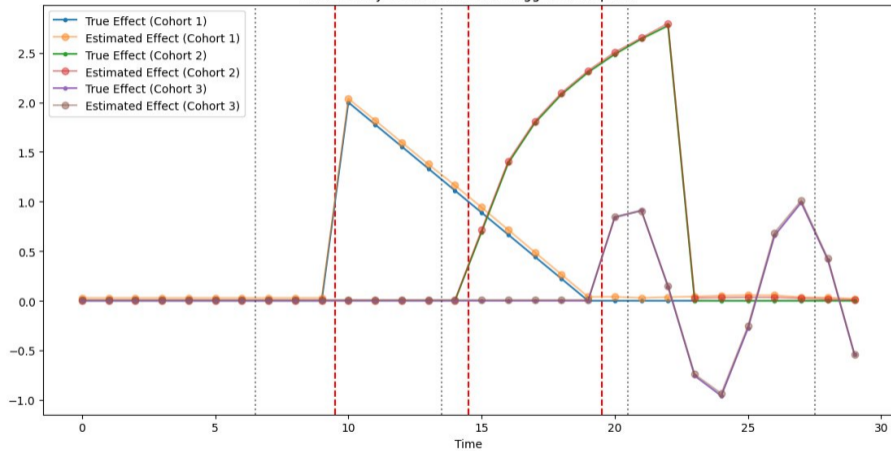


Event Study figures for the four scenarios

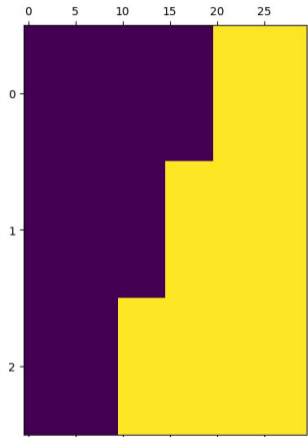
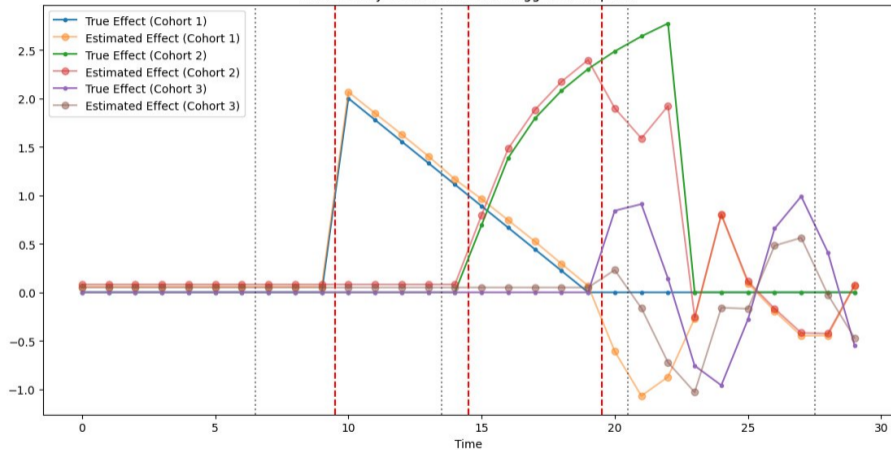




Event Study estimates under staggered adoption



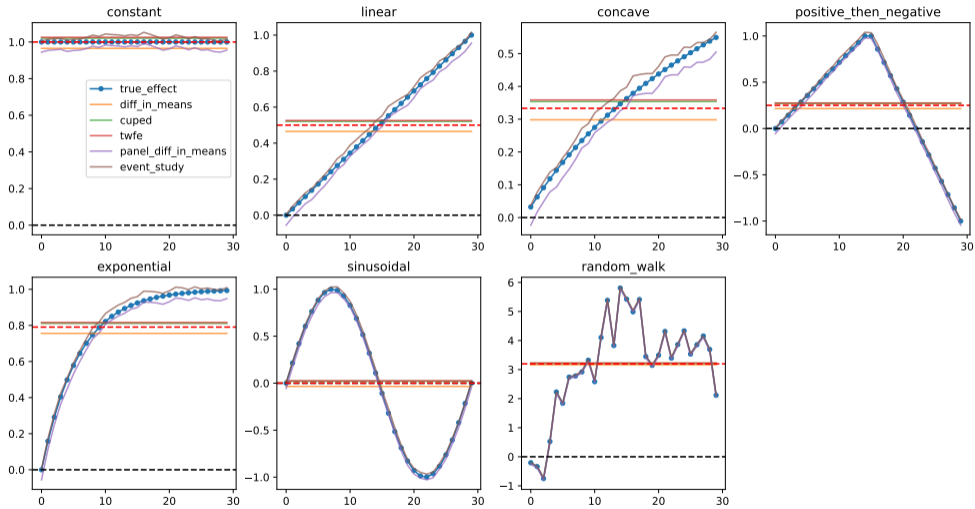
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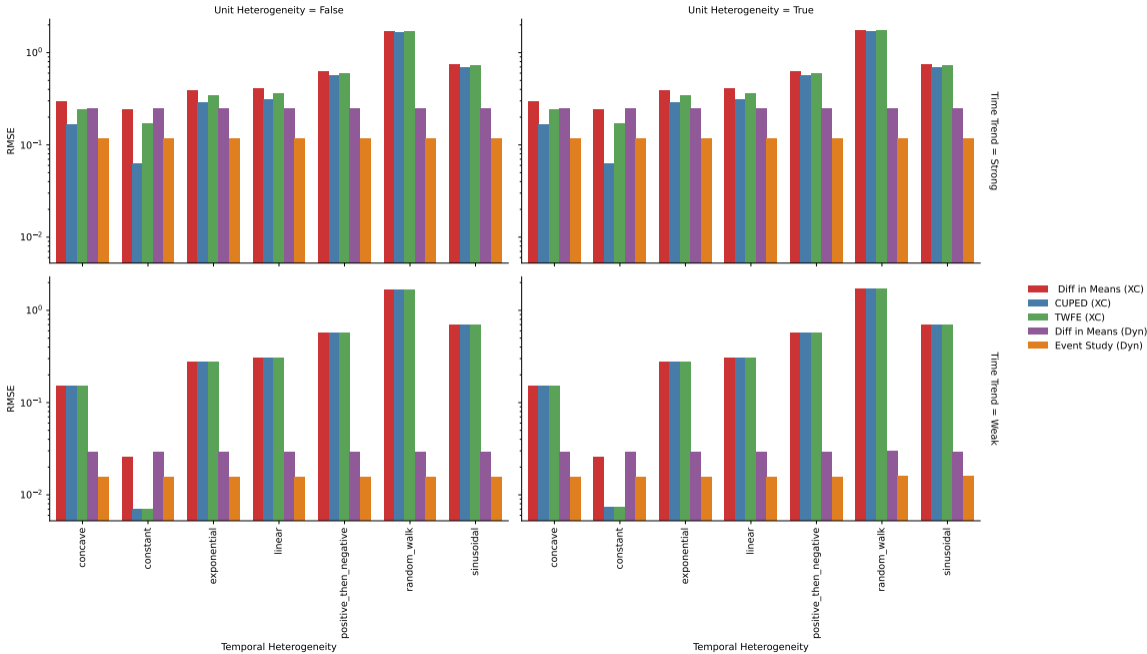
Numerical Experiments



Estimation Accuracy for Different forms of Temporal Heterogeneity



RMSE By (Unit | Time) Heterogeneity and Misspecification



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Runtime Comparisons

$1000 \times \{14, 28, 42, 140, 280, 420\}$ units, $\{14, 28, 42\}$ days.

Observations	Units	Periods	duckreg	pyfixest	pyfixest compressed	statsmodels
14K	1K	14	0.03	0.18	0.17	2.75
28K	1K	28	0.03	0.19	0.18	6.27
42K	1K	42	0.04	0.20	0.21	9.43
140K	10K	14	0.05	0.22	0.26	x
280K	10K	28	0.06	0.26	0.36	x
420K	10K	42	0.04	0.31	0.47	x
1M	100K	14	0.07	0.64	1.27	x
3M	100K	28	0.21	1.00	2.28	x
4M	100K	42	0.24	1.41	3.43	x
14M	1M	14	1.03	4.88	22.12	x
28M	1M	28	3.41	8.07	62.07	x
42M	1M	42	10.92	13.60	117.63	x
140M	10M	14	19.70	123.88	x	x
Mundlak			✓	-	✓	-
Compression			✓	-	✓	-
Out-of-Memory			✓	-	-	-

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Extension: imbalanced panels, descriptive analysis - AKM Regressions

$$\underbrace{Y_{it}}_{\text{Stickiness}} \sim \underbrace{\alpha_i}_{\text{Member FE}} + \underbrace{\psi_{\mathbf{J}(i,t)}}_{\text{Title-FE}} + \underbrace{x'_{it}\beta}_{\text{Covariates}} + \varepsilon_{it}$$



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Naive estimator has problems, fixes use Jackknife (Kline et al 2020).

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Naive estimator has problems, fixes use Jackknife (Kline et al 2020). Variance components suggest different catalogue / recommendation strategies.

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Extensions, Conclusion

- Methods + Software for compressed, out-of-memory computation of commonly used least-squares panel data estimators
 - duckreg : powered by duckDB - out of memory
 - pyfixest: general purpose in-memory regression package, compression in Polars
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 - `duckreg` : powered by `duckDB` - out of memory
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 - Applicable to both scalable estimation of event studies in experiments (focus of this talk), and difference-in-differences methods in observational settings
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 - Extensions / Links
 - Extension to GLMs: Lumley (2018) - estimate MLE on subsample, perform one-step Fisher-scoring update
 - Sketching methods - randomized linear algebra tricks to approx $X'X$ 'well': Mahoney (2013), Pilanci et al (2018), Dobriban et al (2023)
 - Other ideas?
-