## **Heterogeneous Causal Effects with Machine Learning**

Apoorva Lal

July 24, 2022

Stanford

#### **Heterogeneous Treatment Effects - Setup**

- For i.i.d. observations  $i \in \{1,..,N\}$ , we observe  $\{Y_i, \mathbf{X}_i, W_i\}^N_i$  where:
	- *Y<sup>i</sup> ∈* R is the **outcome**
	- $W_i \in \{0, \ldots, K\}$  is the **treatment assignment**
	- $\textbf{v} \cdot \textbf{X}_i \in \mathbb{R}^k$  is the **feature vector**
- $\bullet\;$  We posit the existence of **potential outcomes**  $Y^0,\ldots,Y^k$  for each unit. Append them into a 'science table' that is  $N \times K$ .

#### **Heterogeneous Treatment Effects - Setup**

- For i.i.d. observations  $i \in \{1,..,N\}$ , we observe  $\{Y_i, \mathbf{X}_i, W_i\}^N_i$  where:
	- *Y<sup>i</sup> ∈* R is the **outcome**
	- $W_i \in \{0, \ldots, K\}$  is the **treatment assignment**
	- $\textbf{v} \cdot \textbf{X}_i \in \mathbb{R}^k$  is the **feature vector**
- $\bullet\;$  We posit the existence of **potential outcomes**  $Y^0,\ldots,Y^k$  for each unit. Append them into a 'science table' that is  $N \times K$ .
- Treatment effects (*estimands*) are defined as functions of *potential outcomes*, and since (*K −* 1)*/K* of them are unobserved, we need assumptions to use *estimators* to compute them using data.

-  $\,$  *Causal Consistency / SUTVA* :  $Y_i = \sum_k \mathbb{1}_{W_i = k} Y_i^k.$  What does this assume?

- $\,$  *Causal Consistency / SUTVA* :  $Y_i = \sum_k \mathbb{1}_{W_i = k} Y_i^k.$  What does this assume?
	- *i*'s outcome is only affected by *i*'s treatment status. This may not be the case in many settings, e.g. with peer effects/interference/spillovers/contagion.
	- In such settings, the potential outcomes are indexed by  $Y^{\mathbf{W}}$ . In the extreme case of unrestricted interference, the 'science table' has width  $K<sup>n</sup>$ . Need new assumptions / different estimands.

- $\,$  *Causal Consistency / SUTVA* :  $Y_i = \sum_k \mathbb{1}_{W_i = k} Y_i^k.$  What does this assume?
	- *i*'s outcome is only affected by *i*'s treatment status. This may not be the case in many settings, e.g. with peer effects/interference/spillovers/contagion.
	- In such settings, the potential outcomes are indexed by  $Y^{\mathbf{W}}$ . In the extreme case of unrestricted interference, the 'science table' has width  $K<sup>n</sup>$ . Need new assumptions / different estimands.
- *Unconfoundedness*: *Y* 1 *, Y* <sup>0</sup> *⊥⊥ Wi|***X***i*. Treatment is as good as random given covariates.
- *Overlap*:  $0 < \pi^w(\mathbf{X}) < 1$ . Each unit has positive probability of treatment.

- $\,$  *Causal Consistency / SUTVA* :  $Y_i = \sum_k \mathbb{1}_{W_i = k} Y_i^k.$  What does this assume?
	- *i*'s outcome is only affected by *i*'s treatment status. This may not be the case in many settings, e.g. with peer effects/interference/spillovers/contagion.
	- In such settings, the potential outcomes are indexed by  $Y^{\mathbf{W}}$ . In the extreme case of unrestricted interference, the 'science table' has width  $K<sup>n</sup>$ . Need new assumptions / different estimands.
- *Unconfoundedness*: *Y* 1 *, Y* <sup>0</sup> *⊥⊥ Wi|***X***i*. Treatment is as good as random given covariates.
- *Overlap*:  $0 < \pi^w(\mathbf{X}) < 1$ . Each unit has positive probability of treatment.

Then, the *Counterfactual mean* is non-parametrically identified, as are causal contrasts. AIPW estimator:

$$
\widehat{\Gamma}_i^{(w)} = \underbrace{\widehat{\mu}^w(\mathbf{X})}_{\text{outcome Model}} + \underbrace{\frac{1}{\widehat{\pi}^w(\mathbf{X})}}_{\text{(Inv) Property}} (Y_i - \widehat{\mu}^w(\mathbf{X}))
$$

- $\,$  *Causal Consistency / SUTVA* :  $Y_i = \sum_k \mathbb{1}_{W_i = k} Y_i^k.$  What does this assume?
	- *i*'s outcome is only affected by *i*'s treatment status. This may not be the case in many settings, e.g. with peer effects/interference/spillovers/contagion.
	- In such settings, the potential outcomes are indexed by  $Y^{\mathbf{W}}$ . In the extreme case of unrestricted interference, the 'science table' has width  $K<sup>n</sup>$ . Need new assumptions / different estimands.
- *Unconfoundedness*: *Y* 1 *, Y* <sup>0</sup> *⊥⊥ Wi|***X***i*. Treatment is as good as random given covariates.
- *Overlap*:  $0 < \pi^w(\mathbf{X}) < 1$ . Each unit has positive probability of treatment.

Then, the *Counterfactual mean* is non-parametrically identified, as are causal contrasts. AIPW estimator:

$$
\widehat{\Gamma}^{(w)}_i = \underbrace{\widehat{\mu}^w(\mathbf{X})}_{\text{outcome Model}} + \underbrace{\frac{1}{\widehat{\pi}^w(\mathbf{X})}}_{\text{(Inv) Property}} (Y_i - \widehat{\mu}^w(\mathbf{X}))
$$

- $\hat{\mu}^w(\cdot), \hat{\pi}^w(\cdot)$  *are <i>nuisance functions* (potentially) high-dim quantities incidental to low-dim target (marginal mean, causal contrast).
- All nuisance functions are henceforth *cross-fit*
- Focus (w.log) on binary treatment case
- We are interested in the **Conditional Average Treatment Effect (CATE)**:

$$
\tau(\mathbf{X}) = E[Y^{(1)} - Y^{(0)} | \mathbf{X} = \mathbf{x}]
$$

- This is a *function*, not a number, so we may want to summarise
	- projecting imputed effects linearly on covariates (BLP)
	- binning estimates (GATE)
- $Y_i = \beta_0 + \beta_1 W_i + \beta_2 X_i + \beta_3 W_i X_i + \epsilon_i$ 
	- Implicit outcome models:  $Y_i^0 = \beta_2 X_i$ ,  $Y_i^1 = Y_i^0 + \beta_1 + \beta_3 X_i$
- $\widehat{\text{CATE}}_X = \hat{\beta}_1 + \hat{\beta}_3 X_i$
- Why do we need machine learning / regularization to do this?
- $Y_i = \beta_0 + \beta_1 W_i + \beta_2 X_i + \beta_3 W_i X_i + \epsilon_i$ 
	- Implicit outcome models:  $Y_i^0 = \beta_2 X_i$ ,  $Y_i^1 = Y_i^0 + \beta_1 + \beta_3 X_i$
- $\widehat{\text{CATE}}_X = \hat{\beta}_1 + \hat{\beta}_3 X_i$
- Why do we need machine learning / regularization to do this?
	- **Overfitting**: We know that in general, when *k ≈ N*, traditional OLS methods will badly overfit
	- **Unknown Functional Form**: The analyst does not know what the underlying heterogeneity looks like
	- **fishing**: Why should the reader believe that this specification fell from the sky?

#### **T-Learner**

- fits separate models on the treated and controls.
- $\cdot$  Learn  $\hat{\mu}_{(0)}(x)$  by predicting  $Y_i$  from  $X_i$  on the subset of observations with  $W_i=0.$
- $\cdot$  Learn  $\hat{\mu}_{(1)}(x)$  by predicting  $Y_i$  from  $X_i$  on the subset of observations with  $W_i=1.$
- Report  $\hat{\tau}(x) = \hat{\mu}_{(1)}(x) \hat{\mu}_{(0)}(x)$ .

#### **T-Learner**

- fits separate models on the treated and controls.
- $\cdot$  Learn  $\hat{\mu}_{(0)}(x)$  by predicting  $Y_i$  from  $X_i$  on the subset of observations with  $W_i=0.$
- $\cdot$  Learn  $\hat{\mu}_{(1)}(x)$  by predicting  $Y_i$  from  $X_i$  on the subset of observations with  $W_i=1.$
- Report  $\hat{\tau}(x) = \hat{\mu}_{(1)}(x) \hat{\mu}_{(0)}(x)$ .

### **S-Learner**

- fits a single model to all the data.
- $\bm\cdot\,$  Learn  $\hat\mu(z)$  by predicting  $Y_i$  from  $Z_i:=(X_i,W_i)$  on all the data.
- Report  $\hat{\tau}(x) = \hat{\mu}((x, 1)) \hat{\mu}((x, 0)).$

#### **They were bad: Regularization Bias**



- Differential shrinkage across treatment levels leads to 'hallucinated' heterogeneity
- Problem is generic for any regression learner. Need some kind of 'joint' modelling for potential outcomes.

# **Sidestepping Regularisation Bias: Tailored Neural-net achitecture**

Dragonnet, Tarnet, etc.

$$
\hat{\theta} = \underset{\theta}{\operatorname{argmin}} \,\hat{R}(\theta; \mathbf{X}) \text{ where}
$$
\n
$$
\hat{R}(\theta; \mathbf{X}) = \frac{1}{n} \sum_{i=1}^{n} ((Q^{nn}(w_i, \mathbf{X}_i, \theta) - y_i)^2 + \alpha \text{CrossEntropy}(g^{nn}(\mathbf{X}_i; \theta), w_i))
$$
\n
$$
\hat{\theta} = \underset{t=0}{\operatorname{argmin}} \,\hat{R}(\theta; \mathbf{X}) \text{ where}
$$

https://arxiv.org/pdf/1906.02120.pdf

#### **Sidestepping Regularisation Bias: X, R Learners**

#### **X-Learner**

#### **R-Learner**

- Fit  $\hat{\mu}^{(0)}(x), \hat{\mu}^{(1)}(x)$  using nonparametric regression
- Define pseudo-effects  $\widetilde{D}_i^1 := Y_i - \widehat{\mu}^{(0)}(\mathbf{X}_i)$  and<br>use them to fit  $\widehat{\approx}^1(\mathbf{X}_i)$  on use them to fit  $\widehat{\tau}^1(\mathbf{X}_i)$  on *{i* : *W<sup>i</sup>* = 1*}*
- Define pseudo-effects  $\widetilde{D}_i^0 := \widehat{\mu}^{(1)}(\mathbf{X}_i) - Y_i$  and<br>use them to  $\widetilde{\mathbf{r}}$  *i*  $\widehat{\alpha}^{0}(\mathbf{v}_i)$  on use them to fit  $\widehat{\tau}^0(\mathbf{X}_i)$  on *{i* : *W<sup>i</sup>* = 0*}*
- Aggregate them as  $\widehat{\tau}(x) = (1 \hat{\pi}(x)\hat{\tau}^1(\mathbf{x}) + \hat{\pi}(x)\hat{\tau}^0(\mathbf{x})$

https://arxiv.org/abs/1706.03461

• Minimise Robinson (R) Loss  $\widehat{\tau} = \operatorname*{argmin}_{\tau}$  $\left\{\widehat{L}_n(\tau(\cdot))+\Lambda_n(\tau(\cdot))\right\}$  $\widehat{L}(\tau(\cdot)) = \frac{1}{n}$  $\sum_{n=1}^{n}$  $\sum_{i=1}^{n} ((Y_i - \widehat{\mu}(\mathbf{X}_i)) -$ 

$$
\left(W_i-\widehat{\pi}(\mathbf{X}_i)\right)\tau(\mathbf{X}_i))^2
$$

\n- IOW, Regress pseudo outcome 
$$
\frac{Y - \mu(\mathbf{X})}{W - \hat{\pi}(X)}
$$
 on covariates  $\psi(\mathbf{X}_i)$
\n- weights  $(W - \hat{\pi}(\mathbf{X}))^2$
\n

https://arxiv.org/abs/1712.04912

**DR-Learner**

- Construct pseudo-outcomes  $\widehat{\varphi}(Z) := \widehat{\Gamma}_i^1 - \widehat{\Gamma}_i^0$  using AIPW score function
- Regress it on covariates  $\psi(\mathbf{X}_i)$

### https://arxiv.org/abs/2004.14497

## **In action: RCT, Confounding**

• Simulation + Implementation

## **Experiment**





# **Summary of Generic Approaches [Knaus et al 2021]**



•  $D_i = W_i \in \{0, 1\}$ 

• 
$$
T_i = 2D_i - 1 \in \{-1, 1\}
$$

• 
$$
Y_{\text{IPW}}^* = \frac{W_i - \pi(\mathbf{X}_i)}{\pi(\mathbf{X}_i)(1 - \pi(\mathbf{X}_i))}
$$

• 
$$
Y_{DR}^* = \widehat{\Gamma}_i^1 - \widehat{\Gamma}_i^0
$$

• All problems solve weighted least squares

$$
\min_{\tau} \left( \frac{1}{n} \sum_{i=1}^{n} w_i (Y_i^* - \tau(\mathbf{X}_i))^2 \right)
$$

https://arxiv.org/abs/1810.13237

## **Evaluating HTE Estimators**

#### **Stratification**

- Since Het-FX estimators produce estimates of  $\widehat{\tau}_i$ , a gut-check for how well this works is to then stratify on  $\widehat{\tau}_i$  (say,  $J$  bins), and compute  $\widehat{\text{ATE}}^j$  in each bin using say AIPW
- If  $\widehat{\text{ATE}}^j$ s are sorted along their bin indices, this increases confidence that  $\widehat{\tau}_i$ s aren't all noise

#### **Best linear predictor method**

- Create synthetic predictors  $C_i = \overline{\tau}(W_i - \widehat{\pi}^{-i}(\mathbf{X}_i))$  and  $D = (\hat{\tau}^{-i}(\mathbf{X}_i) - \overline{\tau})(W_i - \hat{\pi}(\mathbf{X}_i))$
- Regress  $Y_i \widehat{\mu}^{-i}(\mathbf{X}_i) \sim \alpha C_i + \beta D_i$
- *α ≈* 1 indicates quality of ATE
- *β ≈* 1 indicates quality of CATE estimates (p.value is an omnibus test of heterogeneity fit by  $\widehat{\tau}_i$ )
- https://datascience.quantecon.org/applications/heterogeneity.html
- https://grf-labs.github.io/grf/articles/diagnostics.html

## **Rank Average Treatment Effects (RATE)**

- Define a targeting rule *S*(**X***i*) which may be based on  $\widehat{\tau}$ s, risk scores, costs (typical  $S$  is simply  $\tau_i$ )
- Define the Targeting Operator Characteristic (TOC) given distribution  $\mathbb{F}(S(\mathbf{X}_i))$  and  $q \in (0,1]$

$$
\text{TOC} = \mathbb{E}\left[Y_i^1 - Y_i^0 | S(\mathbf{X}_i) \ge \mathbb{F}_{S(\mathbf{X}_i)}^{-1}(1-q)\right] - \mathbb{E}\left[Y_i^1 - Y_i^0\right]
$$

• This is largest for small *q*s and decays down to the ATE. If RATE *≈* 0, not much gain from prioritisation

https://grf-labs.github.io/grf/articles/rate.html



- Model-free estimation of ATE and friends largely settled : DML
- In contrast, CATE estimation is a very active area of research
- No silver bullets; good estimators typically depend on substantive knowledge of DGP [Smooth v sparse, etc]
	- prefer estimators that don't bake in function form (e.g. X,R,DR)
- Also prefer estimators that account for confounding (even in RCTs) because of incidental imbalance
- What to do with estimates? Optimal assignment policy learning, AUTOC, etc.